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#### **ABSTRACT**

Cross-validated classification accuracies were compared under assumptions of equal and varying degrees of unequal prior probabilities of group membership for 24 bootstrap and 48 simulated data sets. The data sets varied in sample size, number of predictors, relative group size, and degree of group separation. Total-group hit rates were used to compare the relative accuracies across six assumptions about prior probabilities. Contrary to expectations, use of population priors did not always yield the highest hit rate. When group sizes were similar, equal priors yielded greater classification accuracy than sample estimated priors. Results suggest that, when group sizes are similar, use of unequal priors may lead to a decrement in classification accuracy, even with knowledge of population priors. (Contains 5 tables and 13 references.) (Author)

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# Assuming Equal vs. Unequal Prior Probabilities of Group Membership in Discriminant Analysis: Effect on Predictive Accuracy

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Paper presented at the annual meeting of the American Educational Research Association, April 1995, San Francisco, CA

Assuming Equal vs. Unequal Prior Probabilities of Group Membership in Discriminant Analysis: Effect on Predictive Accuracy

ABSTRACT. Cross-validated classification accuracies were compared under assumptions of equal and varying degrees of unequal prior probabilities of group membership for 24 bootstrap and 43 simulated data sets. The data sets varied in sample size, number of predictors, relative group size, and degree of group separation. Total-group hit rates were used to compare the relative accuracies across six assumptions about prior probabilities. Contrary to expectations, use of population priors did not always yield the highest hit rate. When group sizes were similar, equal priors yielded greater classification accuracy than sample estimated priors. Results suggest that, when group sizes are similar, use of unequal priors may lead to a decrement in classification accuracy, even with knowledge of population priors.

Theoretically, the assumption of unequal prior probabilities should lead to higher cross-validated classification accuracy as the difference between population group sizes increases. "Where we might tend to oversupply small groups and undersupply large ones by using resemblance as the sole basis for classification we introduce a corrective effect by taking prior probabilities of group membership into account" (Tatsuoka, 1988, p. 360). Consistent with this expectation, Rudolph and Karson (1988) found that estimated error rates using population priors were consistently lower than estimated error rates using equal priors.

Although classification accuracy should increase with knowledge and use of population group sizes, these values are rarely known. Consequently, sample estimated values are generally used. However, the use of sample estimated values may be unwise. Huberty (1994, p. 65) argues that "priors should *not* correspond to the relative sample sizes unless ... a proportional sampling plan was utilized." Of course, proportional sampling presumes knowledge of population priors. Lindeman, Merenda, and Gold (1980, p. 211) point out that, "in most practical applications, the values of the prior probabilities are not known with sufficient accuracy to justify their use." Hence, these researchers have urged caution in using anything but equal prior probabilities of group membership for classification.

The purpose of this study was to compare assumptions of equal versus varying degrees of unequal prior probabilities of group membership on cross-validated classification accuracy. The goal was to get some idea of the degree of difference in accuracy we might expect on application of these assumptions in practical classification problems. Implicit in this goal is the question of whether the increment



that may be afforded by assuming unequal priors is worth the risk when population priors are unknown.

#### Method

Cross-validated classification accuracies were compared under a variety of bootstrap and simulated data conditions (different sample sizes, predictor counts, relative group sizes, and prior probability assumptions) for the two-group classification problem. A total of 24 bootstrap and 48 simulated data sets were considered for each of six assumptions about prior probabilities of group membership:

- (1) sample n / sample N (Sample condition);
- (2) 1 / number of groups (Equal condition);
- (3) population n / population N (Pop+0 condition);
- (4) group size for smaller group is 15% larger (Pop+.15 condition);
- (5) group size for smaller group is 30% larger (Pop+.30 condition); and
- (6) group size for smaller group is 45% larger (Pop+.45 condition).

The bootstrap data sets were obtained from 24 real data sets used in a prior classification methodology study (Morris & Huberty, 1987). No pathological distributional problems are known in any of the data sets; it is expected that they are much as one would find in typical classification studies.

The 48 simulated populations were constructed according to multivariate normal models, with N ranging from 1270 to 2000. The group means and covariance matrices needed for input to the population creation program were obtained from the 24 real data sets mentioned in the previous paragraph. For 24 populations, group sizes were set to 1000. The remaining 24 populations were identical to these, except that group sizes were proportional to the sample sizes found in the real data sets.

The process for creating a population manifesting a specified covariance matrix is described in Morris (1975). The random normal deviates required by this method were created using the "Rectangle-Wedge-Tail" method (Marsaglia, MacLaren, & Bray, 1964), with the required uniform random numbers generated by Park & Miller's (1988) "minimal standard" algorithm. A FORTRAN computer program (modified for 64-bit word MS FORTRAN 5.0) provided by Dolker and Halperin was used for the variable generation.

Classification rules for a randomly selected (with replacement) sample of the desired size were built with adjustments made for each of the six assumptions about prior probabilities. The adjusted classification rules were used to classify the entire



population according to Tatsuoka's (1988) minimum chi square rule. This procedure was repeated 1000 times for the 24 bootstrap data sets and 250 times for each of the 48 simulated populations, and the mean number of total-group correct classifications was used to compare the relative accuracies of the six assumptions.

In order to be more confident in the results of this simulation, and in accord with Knuth's (1969, p. 156) recommendation that "the most prudent policy for a person to follow is to run each Monte Carlo program at least twice using quite different sources of random numbers, before taking the answers of the program seriously," the entire simulation results were replicated. In the replication, Wichman and Hill's (1982) uniform random number generator was used. This algorithm generates uniform random numbers by a triple modulo method. As described in Wilkinson (1987, p. 34 of the DATA module), "each uniform is constructed from three multiplicative congruential generators with prime modulus," using 13579, 12345, and 313 as initial seeds. While there were some small differences in the results of the replication, none were systematic, none were judged of importance, and the implications were the same. These replication results are available on request; the results presented in this paper are from the first random number generation method mentioned.

#### Results

For each of the data sets, Tables 1, 2, and 3 give a short description, an index of group separation ( $\underline{\mathbf{D}}$ ), the number of cases in each group ( $\underline{\mathbf{n}}_1$  and  $\underline{\mathbf{n}}_2$ ), the number of predictor variables ( $\underline{\mathbf{p}}$ ), and a comparison of the cross validated classification performance for each assumption about prior probabilities. Tables 2 and 3 also include an index of disproportionality ( $\underline{\mathbf{I}}$ ), calculated as ( ( $\underline{\mathbf{n}}_{larger}$  /  $\underline{\mathbf{n}}_{smaller}$ ) \* 100). The best performing assumptions are underlined. The difference in performance between underlined and nonunderlined assumptions was considered statistically and practically significant based on subjectively established criteria ( $\alpha$  = .00001 plus a mean difference in hit rates of .002, which represents 4 hits for data sets with 2000 cases). The risk of a Type I error was actually much higher than .00001 due to the large number of significance tests conducted. Although statistical significance was considered less important than practical significance, an overall Hotelling  $T^2$  test, and then pairwise post hoc comparisons (multivariate analog of the Scheffé post hoc test; see Morrison, 1976, p. 147-148 for a description) were used to contrast the classification hit rates for the six assumptions.



# Results of Simulation for Data Sets with Equal Group Sizes (#1-24)

The Equal and Pop+0 assumptions, which yield identical results with equal group sizes, were expected to outperform the other four assumptions in all 24 data sets. As indicated in Table 1, the Equal and Pop+0 conditions were top contenders in all but one data set (#15), and yielded the highest (though not always significantly higher) hit rates in 18 data sets (#5 - 8, 10 - 14, 16 - 24). Thus, these assumptions were the best performers most of the time rather than all of the time, which was somewhat contrary to expectations.

# Insert Table 1 About Here

The Sample assumption was expected to perform less well than the Equal and Pop+0 assumptions due to sampling error inherent in the random sampling process, but was still expected to outperform the three erroneous assumptions (Pop+.15, Pop+.30, Pop+.45). Results were consistent with this expectation. The Sample assumption was a top contender in the same 23 data sets as the Equal and Pop+0 assumptions. Nevertheless, compared to the Equal and Pop+0 assumptions, the Sample assumption yielded lower hit rates (though not significantly, based on  $\alpha = .00001$ ) in 21 of the data sets (4-24).

The rank order of the erroneous assumptions was expected to be Pop+.15, Pop+.30, and Pop+.45 (i.e., from least to most discrepant with actual group sizes). Results were consistent with this expectation. The Pop+.15 condition performed better than the other two erroneous assumptions and worse than the Equal, Pop+0, and Sample assumptions. The Pop+.15 assumption was a top contender in 12 of the 24 data sets, and was the best performer in two data sets (#9, 15). The Pop+.30 assumption significantly outperformed the Pop+.45 assumption in 19 of the 24 data sets (#6 - 24).

Results of Simulation for Data Sets with Group Sizes Proportional to Real Data Set Sizes (#25-48)

The Pop+0 assumption was expected to outperform the other five assumptions in all 24 data sets. As shown in Table 2, the Pop+0 condition was a top contender in all but four data sets (#39 - 42), and yielded the highest (though not always significantly higher) hit rates in 11 data sets (#30 - 34, 38, and 43 - 47). No other



assumption performed as well. Thus, although the Pop+0 was the best performer overall, the results were somewhat contrary to expectations because this assumption did not yield the highest hit rate with every data set.

# Insert Table 2 About Here

The Sample assumption was expected to perform less well than the Pop+0 assumption, again due to sampling error, but to outperform the four other assumptions. Results were consistent with this expectation for the Pop+0, Pop+.15, Pop+.30, and Pop+.45 assumptions. Compared to the Pop+0 assumption, the Sample assumption yielded lower hit rates in 20 of the data sets (#27, 28, 30 - 47), though this difference was statistically significant only for data set #45. The Sample assumption was a top contender in 18 of the 19 data sets for which the Pop+0 assumption was also a top contender (#25 - 37, 39, 43, 44, and 46-48), and had the highest hit rates (though not significantly higher) in 2 data sets (#26 and 48). None of the erroneous assumptions matched this performance. Compared to the Pop+.15 condition, which was the best performing erroneous assumption, the Sample assumption yielded higher hit rates (though not always significantly higher) in 15 data sets (#26, 30 - 34, 36 - 38, 43 - 48).

In the 20 data sets with unequal group sizes, the Equal assumption worked better than the Sample assumption only in data sets with small differences between group sizes (#27, 29, 35 - 37, 40 - 42, 44, 45). In the seven data sets with an index of disproportionality greater than 129 (#30 - 32, 34, 38, 43, 47), the Sample assumption outperformed the Equal assumption. The Sample assumption also outperformed the Equal assumption in three data sets with smaller differences between group sizes (#33, 46, 48).

As with equal group sizes, the rank order of the erroneous assumptions with unequal group sizes was expected to be Pop+.15, Pop+.30, and Pop+.45 (i.e., from least to most discrepant with actual group sizes). Results were consistent with this expectation, parallel to the findings for equal group sizes. The Pop+.15 condition performed better than the Pop+.30 and Pop+.45 assumptions and worse than the Pop+0 and Sample assumptions. The Pop+.15 assumption was a top contender in 14 of the 24 data sets, was the best performer in five data sets (# 29, 39 - 42), and outperformed (though not always significantly) the Pop+.30 assumption in 20 data



sets (#29 - 48). The Pop+.30 assumption outperformed the Pop+.45 assumption, though not always significantly, in 21 of the data sets (#26 and 29 - 48).

# Results for Bootstrap Data Sets (Data Sets 49-72)

Results for the bootstrap data sets were quite similar to results for the simulated data sets. As shown in Table 3, the Pop+0 condition was a top contender in all data sets, and had the highest hit rate in 12 data sets (#54 - 56, 58, 61 - 62, and 67 - 72). Nevertheless, other assumptions yielded higher hit rates (though not always significantly higher) in eight data sets (#51, 53, 57, 59, 60, 64 - 66). Thus, as with the simulated data sets, these results were somewhat contrary to expectations because the Pop+0 assumption did not yield the highest hit rate with every data set.

#### Insert Table 3 About Here

The Sample assumption was a top contender in all but one data set (#70), and had the second highest hit rate (behind Pop+0) in nine data sets (#54 - 56, 58, 61, 62, 67, 71, 72). Compared to the Pop+0 assumption, the Sample assumption yielded lower hit rates in 22 of the data sets (#51 - 72), though this difference was statistically significant only for data set #70. Compared to the Pop+.15 condition, the Sample assumption yielded higher hit rates (though not always significantly higher) in 14 data sets (#52, 54, 55, 56, 58, 61 - 63, 67 - 72). Thus, the performance of the Sample assumption relative to the erroneous assumptions matches what was found in the simulated data sets.

The Equal assumption worked better than the Jample assumption only in data sets with similar group sizes (#51 - 53, 57, 59, 60, 63 - 66, 68 - 70). The Sample assumption outperformed the Equal assumption in the seven data sets with indices of disproportionality greater than 129 (#54 - 56, 58, 62, 67, 71), as well as in two data sets with more similar group sizes (#61, 72). The Pop+.15 condition performed better than the Pop+.30 and Pop+.45 assumptions and worse than the Pop+0 and Sample assumptions. The Pop+.15 assumption was a top contender in all but three data sets (#61, 63, 70), was the best performer in two data sets (57, 65), and outperformed (though not always significantly) the Pop+.30 assumption in 20 data sets (#52, 54 - 72). The Pop+.30 assumption outperformed the Pop+.45 assumption, though not always significantly, in 21 of the data sets (#52 - 67). Again, these bootstrap results were similar to what was found in the simulated data sets.



### Discussion

# Pop+0 vs. Other Assumptions

Although the Pop+0 assumption was the best performer in an absolute sense, its performance relative to the other five assumptions was not as good as predicted. The erroneous assumptions occasionally performed much better than would be expected based on their discrepancy with population sizes. For example, in some data sets with unequal population sizes, the erroneous assumption of equal priors yielded a higher hit rate than the correct assumption about population priors.

At first glance, these results appear to be inconsistent with Rudolph and Karson's (1988) finding of consistently lower error rate estimates using population priors rather than equal priors. This apparent inconsistency may be due to differences in relative population sizes between the two studies. In the Rudolph and Karson study, the population priors were .9 and .1, reflecting a large discrepancy in population sizes. In the present study, equal priors yielded a higher hit rate than population priors only in data sets with similar group sizes. In all data sets with non-trivial differences in group sizes (I greater than 129), use of population priors increased the hit rate over equal priors.

Further support for this explanation of the apparent inconsistency between the two studies comes from a partial replication of the simulation. For data sets with similar group sizes (I less than or equal to 129) in which an erroneous assumption outperformed the Pop+0 assumption, new simulated data sets were created, each with 900 1's and 100 2's. As in the Rudolph and Karson study, the Pop+0 assumption outperformed the erroneous assumptions for every data set. These results are displayed in Table 4. Thus, our findings were consistent with Rudolph and Karson for data sets with dissimilar group sizes.

# Insert Table 4 About Here

Still, it may seem counterintuitive that in <u>any</u> data set, erroneous assumptions about priors could yield higher hit rates than the correct assumption. An explanation for this is related to the differential effectiveness of the two classification rules when different priors are used. Suppose the two classification rules are equally effective in classifying 1's and 2's using equal priors. What happens when the rules are adjusted



for unequal priors? "For groups of unequal sizes that tend to reflect relative population sizes (in an order sense), use of unequal priors will increase the hit rates for the larger groups and decrease the hit rates for the smaller groups" (Huberty, 1994, p. 112). When the increment in hits for the larger group exceeds the decrement in hits for the smaller group, the overall hit rate is higher. However, when the decrement in hits for the smaller group exceeds the increment in hits for the larger group, the overall hit rate is lower.

Consider data set #41 (Table 2), in which the correct (Pop+0) assumption about priors yields a lower hit rate than three of the incorrect assumptions (Equal, Pop+.15, Pop+.30). For this data set, Table 5 displays the average separate group and total hits for each of the six assumptions about priors. We can see how changes in separate group hits affect the results. Compared to the Equal assumption, for example, the Pop+0 assumption averages 19 more hits for Group 1 but 24 fewer hits for Group 2. Consequently, there are fewer total hits for the correct Pop+0 assumption than for the incorrect assumption of equal priors.

#### Insert Table 5 About Here

# Sample-Estimated Priors vs. Equal Priors

Relative to the assumption of equal priors, the assumption of sample-estimated priors mirrored the Pop+0 pattern. When group sizes were similar, the Equal assumption was generally superior. When group sizes differed by 13% or more, the Sample assumption outperformed the Equal assumption.

Huberty (1994, p. 65) contends that sample estimated priors are inappropriate unless proportional sampling has been used. Results from the current study suggest that, perhaps even with proportional sampling, use of population priors may lead to a decrement in classification accuracy when group sizes are similar. Additional study is needed to confirm this interpretation.



# References

- Dolker, M., & Halperin, S. (1982). Comparing inverse, polar, and rectangle-wedge-tail FORTRAN routines for pseudo-random normal number generation. Educational and Psychological Measurement, 42, 223-226.
  - Huberty, C. J (1994). Applied discriminant analysis. New York: Wiley.
- Knuth, D. E. (1969). The art of computer programming (Vol. 2: Seminumerical algorithms). Reading: MA: Addison-Wesley.
- Lindeman, R. H., Merenda, P. F., & Gold, R. Z. (1980). <u>Introduction to bivariate and multivariate analysis</u>. Glencoe, IL: Scott, Foresman.
- Marsaglia, G., MacLaren, D., & Bray, T. A. (1964). A fast procedure for generating random normal variables. Communications of the ACM, 7, 4-10.
- Morris, J. D. (1975). A computer program to create a population with any desired centroid and covariance matrix. <u>Educational and Psychological Measurement</u>, 35, 707-710.
- Morris, J. D., & Huberty, C. J (1987). Selecting a two-group classification weighting algorithm. <u>Multivariate Behavioral Research</u>, 22, 211-232.
- Morrison, D. F. (1976). <u>Multivariate statistical methods</u>. New York: McGraw-Hill.
- Park, S. K., & Miller, K. W. (1988). Random number generators: Good ones are hard to find. Communications of the ACM, 31, 1192-1201.
- Rudolph, P. M., & Karson, M. (1988). The effect of unequal priors and unequal misclassification costs on MDA. <u>Journal of Applied Statistics</u>. 15, 69-81.
- Tatsuoka, M. M. (1988). <u>Multivariate analysis: Techniques for educational and psychological research</u> (2nd ed.). New York: Macmillan.
- Wichman, B. A., & Hill, I. D. (1982). An efficient and portable pseudorandom number generator. Algorithm AS 183. Applied Statistics, 311, 188-190.
- Wilkinson, L. (1987). <u>SYSTAT: The system for statistics</u>. Evanston, IL: SYSTAT, Inc.



Table 1

Cross Validated Classification Performance (Portion of "Hits") for Six Assumptions About Prior Probabilities for Simulated Data Sets with Equal Group Sizes

.

							Ass	Assumption		
# Data Cat Decoriation		É	ć	c	Sample	2 Equal	3 Pop+0	4 Pop+.15	5 Pop+.30	6 Pop+.45
# Data Set Description			7	2						-
1 Eicher Date Groupe 1 & 3	12 63	1000	1000	4	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
7 Fisher Data - Groups 1 & 2	8.79	0001	1000	4	6666	8666	8666	8666	8666	7666
3 Bishev Data - Groups 1 & 3	5.10	1000	0001	13	.9851	.9851	9851	.9858	.9864	8986
4 Fisher Data - Groups 2 & 3	4.24	000	900	4	.9792	9794	9794	.9804	6086	6086
S Rulon Data - Groups 1 & 3	2.92	1000	1000	4	9206	.9211	.9211	.9190	.9152	.9091
6 Rishey Data - Groups 1 & 2	2.84	1000	000	13	.9028	.9033	.9033	.9031	8006	.8954
Data - Groups 2	2.47	1000	1000	13	.8694	.8700	8700	.8692	.8663	.8602
8 Talent Data - Groups 3 & 5	2.08	1000	900	14	.8314	.8320	.8320	.8311	.8260	8159
9 Demographic # 2 - Body Char	1.86	1000	900	<b>∞</b>	.8140	.8145	.8145	.8152	0608.	.7932
10 Rulon Data - Groups 2 & 3	1.86	900	900	4	.8182	.8188	.8188	.8158	.8062	7891
11 Rulon Data - Groups 1 & 2	1.74	0001	900	4	.8045	8052	.8052	.8046	.7954	.7741
12 Talent Data - Groups 1 & 5	1.73	1000	900	14	7882	7888	.7888	7878	.7809	.7678
13 Demographic # 3 - Body Char	1.35	1000	1000	∞	.7378	7390	7390	.7334	9/1/.	.6928
14 Talent Data - Groups 1 & 3	.92	900	900	14	.6357	.6365	.6365	.6286	.6135	.5924
15 Block Data - Groups 3 & 4	.85	1000	1000	4	9959.	.6621	.6621	.6733	.6590	.6275
16 Block Data - Groups 1 & 2	82	1989	900	4	.6469	.6522	.6522	.6285	.5927	.5563
17 Block Data - Groups 1 & 4	.81	900	9	4	.6348	6392	.6392	.6145	1186.	79487
18 Block Data - Groups 1 & 3	.75	<u>8</u>	8	4	.6255	.6314	.6314	.6037	.26//	.5365
19 Warmcke Data - Groups 1 & 3	<i>L</i> 9.	000	900	01	.5935	.5955	.5955	.5945	.5840	.5657
20 Block Data - Groups 2 & 3	99:	1000	1000	4	.6061	.6095	5609	.5870	.5589	.5335
21 Block Data - Groups 2 & 4	.50	900	900	4	.5657	.5692	5692	.5602	.5433	.5269
22 Demographic # 1 - Body Char	.50	900	99	∞	.5809	.5822	.5822	.5658	.5413	.5206
23 Warncke Data - Groups 1 & 2	.48	1000	900	10	.5602	5604	5604	.5532	.5397	.5253
24 Warncke Data - Groups 2 & 3	.45	1000	1000	10	.5471	.5482	.5482	.5422	.5318	.5203

Note. The best performing assumption(s) are underlined (p < .00001).

Table 2

Cross Validated Classification Performance (Portion of "Hits") for Six Assumptions About Prior Probabilities for Simulated Data Sets with Group Sizes Proportional to Real Data Sets

									Assumption	ption			
# Data Set I	Data Set Description	Q	<b>James</b>	ū	n,	. 57	1 Sample	2 Equal	3 Pop+0	4 Pop+.15	5 Pop+.30	6 Pop+.45	
-	Acsel iption				7								
25 Fisher Data	ta - Groups 1 & 3	12.63	100	1000	1000	4	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	
26 Fisher Data	ī	8.79	8	<u>00</u>	88	4 ;	6666	855	3666		9866	0860	
27 Bisbey Da	_	5.12	90	945	666	5	555	/C85.	5050		0000	0800	
28 Fisher Data	1	4.24	8	8	36	4.	76/6	4/2	47.75		2002	9166	
29 Rulon Data	1	2.92	129	935	726	4 ;	0) 16: 18: 18:	1616	2170		52100	0134	
30 Bisbey Data	ita - Groups 1 & 2	5.8	231	420	972	13	.9168	1016.	11/2		2012	7776	
31 Bisbey Da	Data - Groups 2 & 3	2.41	219	972	444	: 13	8748	8623	46/8		9776	9231	
32 Talent Data	ta - Groups 3 & 5	1.97	285	330	940	4	.8426	.8146	2431		0/00.	1650.	
33 Demograi	္က	1.88	129	942	732	∞ '	.8224	.8182	8231		.8083	2/6/3	
34 Rulon Data	ta - Groups 2 & 3	1.87	141	930	99	4	.8251	.8202	\$203		1618.	.0045	
35 Rulon Da	- Groups 1	1.74	<u>8</u>	820	930	4 :	8002	8023	2017		1940	2677.	
36 Talent Da	Falent Data - Groups 1 & 5	1.72	113	830	940	4.	7848	(8)	1820		10//-	20707	
37 Demogra	Demographic # 3 - Body Char	1.35	<u>\$</u>	994	959	<b>∞</b> ;	(S)	5157	1767		6003	6765	
38 Talent Da	alent Data - Groups 1 & 3	<b>6</b> 8	252	966	396	14		.0455	* 15		6500	2010.	
39 Block Da	Slock Data - Groups 3 & 4	.85	8	8	35	4.	6366	1799.	1700.		0,50	6100	
40 Block Data	ta - Groups 1 & 2	.85	105	900	950	4 -	.644. 7.000	/1co.	. C440.		7386	6132	
41 Block Data	ta - Groups 1 & 4	∞.	105	38	950	4.	.6327	8660.	1/00.		5227	5035	
42 Block Data -	ita - Groups 1 & 3	7.	103	3	975	4 ;	.6214	6779.	7070		1777	1005 1005	
43 Warncke	Warncke Data - Groups 1 & 3	<b>89</b> .	163	975	3	2	.6408	7509.	478		1/00.	100C.	
44 Block Da	Block Data - Groups 2 & 3	99.	105	925	975	4	.6117	.6129	<u> </u>		1955.	6765	
45 Block Da	Block Data - Groups 2 & 4	.50	103	962	886	4	.5655	5699	5704		5424	51.00	
46 Demogra	Demographic # 1 - Body Char	.50	<u>\$</u>	626	994	∞	.5836	.5833	.5845		2040.	2816.	
47 Warncke	Warncke Data - Groups 1 & 2	.48	138	975	705	01	.5801	.5661	.5811		0755.	.5283	
48 Warncke	' (7	4. 4.	118	284	840	0	.5527	.5488	.5521		.5342	8816.	1
40 44 6111011	Dam Courts												1

Note. The best performing assumption(s) are underlined (p < .00001).

<u>П</u>

Table 3

Cross Validated Classification Performance (Portion of "Hits") for Six Assumptions About Prior Probabilities for Bootstrap Data Sets with Group Sizes Proportional to Real Data Sets

Note. The best performing assumption(s) are underlined (p < .00001).



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Table 4

Cross Validated Classification Performance (Portion of "Hits") for Six Assumptions About Prior Probabilities for Simulated Data Sets with 900 1's and 100 2's

								Assumption	ption		
						_	2	3	4	5	9
# Data Set Description	D	I	n	n2	d	Sample	Equal	Pop+0	Pop+.15 I	Pop+.30	Pop+.45
73 Fisher Data - Groups 1 & 2	12.96	8	8	8	4	1.0000	0000	1.0000	1.0000	1.0000	1.0000
74 Rishev Data - Groups 1 & 3	4.74	8	8	81	13	.9873	.9826	9876	.9875	.9874	.9873
75 Fisher Data - Groups 2 & 3	4.02	8	8	8	4	9866	9926.	.9870	6986	6986	2986
	3.06	8	8	8	4	.9628	.9280	.9633	.9628	.9618	.9607
77 Demographic # 2 - Body Char	2.06	8	8	8	<b>∞</b>	.9285	.8462	.9287	.9277	.9266	.9250
78 Rulon Data - Groups 1 & 2	1.93	8	8	90	4	.9301	.8406	.9305	.9295	.9281	.9264
-	1.91	8	906	8	14	.9200	.8341	.9205	.9187	.9167	.9146
_	1.32	8	8	8	∞	.9042	.7459	9044	.9036	.9023	.9004
81 Block Data - Groups 3 & 4	1.23	8	8	8	4	6916	0///	<u> 9167</u>	.9162	.9152	.9141
82 Block Data - Groups 1 & 2	.75	8	8	8	4	8926	.6498	.8928	.8905	8875	.8844
83 Block Data - Groups 1 & 4	.73	8	8	8	4	8933	.6517	.8938	.8916	8888.	.8853
84 Block Data - Groups 1 & 3	89.	8	8	8	4	8948	.6242	.8953	.8937	.8917	.8892
85 Warncke Data - Groups 2 & 3	.43	8	8	8	10	8868	.6764	7988.	.8829	.8787	.8742

Note. The best performing assumption(s) are underlined (p < .00001).

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Table 5

Average Separate Group and Total Hits for Six Assumptions about Priors for Data Set #41

			Ass	umptions		
Average Hits	Sample	Equal	Pop+0	Pop+.15	Pop+.30	Pop+.45
Total	1233.820	1247.644	1242.416	1260.152	1245.284	1195.668
Group 1	633.240	616.400	635.124	526.100	415.496	303.612
Group 2	600.580	631.244	607.292	734.052	829.788	892.056

